Limitations of SRTM, Logan graphical method, and equilibrium analysis for measuring transient dopamine release with $[^{11}\text{C}]$raclopride PET

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Received December 23, 2012; Accepted February 14, 2013; Epub April 9, 2013; Published April 15, 2013

Abstract: Conventional PET methods to estimate $[^{11}\text{C}]$raclopride binding potential ($BP_{ND}$) assume that endogenous dopamine concentration does not change during the scan time. However, this assumption is purposely violated in studies using pharmacological or behavioral stimuli to invoke acute dopamine release. When the assumption of steady-state dopamine is violated, conventional analysis methods may produce biased or even unusable estimates of $BP_{ND}$. To illustrate this problem, we examined the effect of scan duration on $\Delta BP_{ND}$ estimated by three common analysis methods (simplified reference tissue model, Logan graphical reference method, and equilibrium analysis) applied to simulated and experimental single-scan activation studies. The activation – dopamine release – in both the simulated and experimental studies was brief. Simulations showed $\Delta BP_{ND}$ to be highly dependent on the window of data used to determine $BP_{ND}$ in the activation state. A similar pattern was seen in the data from human smoking studies. No such pattern of $\Delta BP_{ND}$ dependence on the window of data used was apparent in simulations where dopamine was held constant. The dependence of $\Delta BP_{ND}$ on the duration of data analyzed illustrates the inability of conventional methods to reliably quantify short-lived increases in endogenous dopamine.

Keywords: Binding potential ($BP_{ND}$), transient dopamine release, model selection, scan duration, smoking, time-invariant

Introduction

Traditional analysis of dynamic $[^{11}\text{C}]$raclopride positron emission tomography (PET) studies includes estimating $[^{11}\text{C}]$raclopride binding potential ($BP_{ND}$) – a static parameter that represents the potential for specific binding of $[^{11}\text{C}]$raclopride to dopamine (DA) D2/D3 receptors in the brain. This is usually accomplished by fitting the dynamic data with compartmental or graphical (linearized) models, such as the simplified reference tissue model (SRTM) [1] or the Logan graphical method [2, 3]. There are no time-varying parameters in any of the conventional models. Thus, a basic assumption of these models is that endogenous dopamine (DA) concentration is at steady-state and does not change during the course of the scan.

However, experiments are often conducted for which the assumption of steady-state DA concentration is purposely violated in order to measure changes in synaptic DA due to a pharmacological or behavioral challenge. In these studies, $BP_{ND}$ is measured at rest (no challenge) and after a challenge. An increase in endogenous DA in the challenge condition competes with the radiotracer for binding at the receptor and results in lower $BP_{ND}$ values relative to the rest condition [4]. In this “competition model”, decreases in $BP_{ND}$ are used as an index of DA release following a wide variety of pharmacological stimuli, including amphetamine [5-7] and nicotine [8-15].

A decrease in $BP_{ND}$ has been shown to be a robust and reliable index of amphetamine-induced DA release. Amphetamine has been shown to produce large increases in DA concentration [16-18] that peak over 600% above baseline [17]. It has also been shown to pro-
voke long-lasting displacement of $[^{11}C]$raclopride (potentially due, in part, to receptor internalization [4, 19]) and greater than 20% decrease in $B_{PD}^{ND}$ [5]. Conventional analysis methods, such as SRTM, Logan graphical method (“Logan plot”), and equilibrium methods, have been able to reproducibly estimate sustained decreases in $B_{PD}^{ND}$ for varying doses of amphetamine administered under a variety of experimental conditions in both humans and animals.

In contrast, it is unclear whether nicotine-induced DA release can be detected by decreases in $[^{11}C]$raclopride $B_{PD}^{ND}$. Based on a study of human subjects smoking 1 cigarette during a break in a scan with bolus-plus-constant-infusion (B/I) of $[^{11}C]$raclopride, Brody et al. reported 30% decrease in $B_{PD}^{ND}$ after smoking in the left ventral caudate [13]. In later studies with larger cohorts and a more advanced scanner, but the same protocol, Brody et al. reported only 8% decrease in $[^{11}C]$raclopride $B_{PD}^{ND}$ [9-12]. Scott et al. reported a similar decrease in $B_{PD}^{ND}$ (~10%) between regular and denicotinized cigarettes smoked while in the scanner [15]. However, only 1% decrease was reported by Montgomery et al. in $[^{11}C]$raclopride $B_{PD}$ in response to 2mg nicotine nasal spray administered during scanning [20], and a study by Barrett et al. requiring subjects to smoke up to 6 cigarettes during the scan measured highly variable changes in $[^{11}C]$raclopride $B_{PD}^{ND}$ (range of +57% and -67% change in $B_{PD}^{ND}$) in the ventral striatum and caudate [8].

We believe that the between-study and within-study variability in nicotine-induced decrease in $B_{PD}^{ND}$ reported in the literature may be due to violations of the assumption that endogenous DA concentration remains at steady-state during the challenge condition. Unlike amphetamine, which is a direct releaser of DA, nicotine only indirectly causes transient increases in DA concentration and transient displacement of $[^{11}C]$raclopride. Microdialysis studies have shown that nicotine, in doses comparable to what a human would self-administer by smoking, produces a brief 125% increase in DA which returns to baseline within 30 min [21]. Nicotine challenges in PET studies may also cause transient decreases in $[^{11}C]$raclopride signal which return to baseline levels just as quickly. Because of amphetamine’s long-lasting effects on DA release and tracer displacement, $[^{11}C]$raclopride reaches a new equilibrium state after challenge and $B_{PD}^{ND}$ can be estimated by conventional models without appreciably violating the assumption that endogenous DA concentration is not changing during the scan. For nicotine challenges, DA concentration and tracer binding are continuously changing and $[^{11}C]$raclopride does not reach a new equilibrium state. In this scenario, specific binding of $[^{11}C]$raclopride will not be constant during the challenge scan and any attempt to estimate $B_{PD}^{ND}$ with conventional methods will produce an estimate which is a weighted-average of the specific tracer binding as it changes over the course of the challenge scan [22]. This could result in a reduced measure of nicotine-induced decrease in $B_{PD}^{ND}$, which, coupled with a small effect size and inter-subject variability, could easily obscure detection of a significant decrease in $B_{PD}^{ND}$ due to DA release from nicotine.

To investigate if inappropriate application of conventional modeling approaches could obscure detection of nicotine-induced DA release, we examined the effect of scan duration on decrease in $B_{PD}^{ND}$ estimated by SRTM, Logan plot, and equilibrium analysis. We applied these analyses to simulated single-scan activation studies and human smoking studies. An observed dependence of $\Delta B_{PD}^{ND}$ on the duration of data analyzed suggests that conventional methods are not well-suited to reliable quantification of short-lived increases in endogenous DA.

**Materials and methods**

**Simulations**

Noiseless and noisy striatal $[^{11}C]$raclopride time activity curves (TACs) with varying amounts of DA release were simulated using an enhanced tracer kinetic model (Figure 1) [23]. This model includes state variables for free and bound dopamine and it includes rate constants for the tracer as well as binding and dissociation of DA to and from the receptor. TACs based on a B/I of tracer were simulated for a scan duration of 120 min consisting of 1-min frames.

For all simulations, striatal parameters were set as $K_f = 0.0918 \text{ mL/min/g}$, $k_{on} = 0.4484 \text{ min}^{-1}$, $k_{off} = 0.0282 \text{ mL/(pmol \cdot min)}$, $B_{max} = 44 \text{ pmol/mL}$, and $F_r = 0.04 \text{ mL/mL}$. 
DA binding parameters (association and dissociation rate) were set as $k_{on}^{DA} = 0.25 \text{ mL/(pmol} \cdot \text{min)}$ and $k_{off}^{DA} = 25 \text{ min}^{-1}$ [23]. A noiseless cerebellum TAC was also simulated using the same model by setting $k_{on}$ and $k_{off}$ to 0. All simulations were implemented in MATLAB (R2006a, MathWorks, Inc., Natick, MA) using modeling functions provided by a library of COMKAT [25].

Dopamine release: Free DA ($F_{DA}$) was modeled as a gamma variate plus a constant baseline with the start of DA release, $t_d$, set at 40 min post tracer injection:

$$F_{DA}(t) = DA_{basal} + \gamma(t - t_d) e^{-\lambda t}$$ for $t \geq t_d \quad (1)$$

Baseline DA level ($DA_{basal}$) was set to 100 pmol/mL [26]. Free DA curves with a fixed peak height but increasing area-under-the-curve (AUC) were created by varying $\alpha$, $\beta$, and $\gamma$ parameters which specify the shape and magnitude of a gamma variate. Five different DA signals were created: $[\alpha, \beta, \gamma] = [2, 2, 800], [2, 1, 200], [2, 0.5, 50], [2, 0.25, 12.5], [2, 0.15, 4.5]$ (Figure 2). All DA signals were designed to represent plausible cases of transient DA release (defined as returning to baseline during the course of the simulated scan). Total DA released was measured as the area under $F_{DA}(t) - DA_{basal}$ from $t_d$ to the end of the scan (e.g., 40 min to 120 min post injection) and ranged from 190-2665 (pmol·min)/mL. Noiseless and noisy Rest TACs with no DA perturbation were simulated by setting $\gamma$ to 0.

Noise: Monte-Carlo simulations were used to produce 100 noisy PET TACs for each DA signal. Additive Gaussian noise proportional to the imaged PET activity was added to the noiseless, simulated curves [27, 28]:

$$ROI(t) = ROI(t) + SD(t) \times G(0,1)$$

where $ROI(t)$ is the noiseless, decay-corrected TAC, $G(0,1)$ is a pseudo-random number from a Gaussian distribution with 0 mean and standard deviation of 1, and $SD(t)$ is the standard deviation of the noise:

$$SD(t) = \sqrt{c \times ROI(t) \times e^{-\lambda t}} \times e^{-\lambda t} \quad (3)$$

In this noise model, $c$ is the proportionality constant which determines the noise level, $\lambda$ is the decay constant for C-11, and $\Delta t$ is the simulated frame length. The term $e^{-\lambda t}$ applies decay to the decay-corrected $ROI(t)$ while the term $e^{-\lambda t}$ decay-corrects the standard deviation. The constant $c$ was set to 0.002 to produce a noise level similar to that seen in regional-level TACs extracted from the ventral striatum in $[^{11}C]$ raclopride studies completed on the High Resolution Research Tomograph (HRRT; Siemens/CTI, Knoxville, TN, USA).

Human experiments

Subjects: Healthy smokers ($n = 4$, 2M/2F) were recruited as subjects for a study investigating the effect of cigarette smoking and the temporal pattern of DA release. Subjects were screened to determine medical history and personal and familial psychiatric and smoking history. They were excluded from the study if they met criteria for any Structured Clinical Interview for DSM-IV Axis 1 Disorders except Nicotine Dependence. Severity of nicotine dependence was assessed by the Fagerström Test for Nicotine Dependence (FTND). Only subjects with an FTND score $> 3$ were included in this study. Smoking status was confirmed by breath CO levels $> 12$ ppm and urine cotinine levels $>
150 ng/mL at intake. Smokers smoked 15 ± 6 cig/day for 13 ± 15 yr. Written informed consent was obtained from all subjects before the study but only after complete explanation of the study procedures. All study procedures were approved by the Yale University Human Investigations Committee (HIC).

**Experimental design:** Prior to PET imaging, each subject received a 3T structural MRI using a 3D fast spoiled grass (FSPGR) MR pulse acquisition sequence with an IR prep of 300 ms (TE = 3.3 ms, flip angle = 17 degrees; slice thickness = 1.2 mm). Subjects received up to two [11C]raclopride scans (rest and smoking) on separate days. Only the data from the smoking scan were used in these analyses. [11C]raclopride was synthesized as reported previously [29]. Coincident with a B/I of [11C]raclopride (685 ± 55 MBq, mean ± standard deviation), list-mode data were acquired for 120 min on the HRRT. A transmission scan was conducted immediately prior to the dynamic scan for attenuation correction. The Vicra optical tracking system (Vicra, NDI Systems, Waterloo, Canada) was used to measure and record subject head movement during each scan. The Vicra records subject head motion in quadrants at a rate of 20Hz from a rigid tool with reflective spheres attached to each subject’s head via a swim cap.

**Smoking challenge:** During the smoking scan, 3 of 4 subjects were asked to smoke 2 consecutive cigarettes, beginning 45 min after the initial bolus of [11C]raclopride. One female subject was asked to smoke two cigarettes ~30 min apart. The first cigarette was at 53 min; the second was cigarette at 87 min after time-of-injection (TOI).

All subjects were allowed to smoke their own brand of cigarette. Immediately prior to the start of smoking, a study investigator informed the subject that they were about to start smoking. The investigator handed the subject a cigarette, which the investigator lit with a long-reach lighter once the subject had placed the cigarette in his/her mouth. Subjects were instructed to smoke at their typical pace while remaining as still as possible. A small basin was provided at the subject’s chest as an ashtray. An investigator monitored the smoking session, and when subjects finished smoking each cigarette, they dropped it into the basin which the study investigator removed immediately to extinguish the cigarette. Each subject was asked if he/she felt they could smoke a second cigarette (either immediately following the first cigarette or 30 min later). All subjects accepted and smoked a second cigarette. A freestanding air exhaust and filtration system (Movex Inc, Northampton, PA) that had been previously approved by the Yale Environmental Health and Safety Office was used to capture and filter second-hand smoke.

**Data processing:** Dynamic scan data were reconstructed iteratively with all corrections, including motion using the information collected by the Vicra, by the MOLAR algorithm [30] with the following frame timing: 40 x 3 min. Frame-by-frame motion correction was also performed by registering each frame to an early summed (0 – 9 min) image using a six-parameter mutual information algorithm (FLIRT, FSL 3.2, Analysis Group, FMRIB, Oxford, UK). A summed (0 – 9 min) motion-corrected image was then registered to each subject’s MR, which was warped to an AAL MR template. Anatomically-based striatal regions of interest (ROIs) were defined on the template in the left and right dorsal caudate, left and right dorsal putamen, and left and right ventral striatum [31] (based on Martinez et al. and Mawlawi et al. [6, 32]). TACs were extracted from these regions as well as from an AAL-defined cerebellum (reference region).

**Estimating BP:**

\[
\Delta B_{PD} = \frac{B_{PD} \text{Baseline} - B_{PD} \text{Activation}}{B_{PD} \text{Baseline}} \times 100
\]

Because we focused on a single-scan study design, \(B_{PD} \text{Baseline}\) was estimated by fitting 0-35 min of data (simulated or experimental) with SRTM and Logan plot (0 min represents time of
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Previous studies have indicated that cues for reward may contribute to DA release [35]. Subjects in human experiments were aware that they would be smoking 45 min after the start of the scan, and study investigators entered the scanning suite up to 5 min before the start of the smoking session. To ensure that the potential cue of study investigators entering the scanning suite did not induce DA release (i.e., expectation prior to smoking) which could bias the $BP_{ND}^{\text{Baseline}}$ value, only 35 min of data were used to calculate $BP_{ND}^{\text{Baseline}}$. For EQ, $BP_{ND}^{\text{Baseline}}$ was taken as the average $BP_{ND}$ from 30-40 min. To test the effect of activation scan duration on $\Delta BP_{EQ}$, $BP_{ND}^{\text{Activation}}$ was estimated from increasingly wider data windows by fitting 0 - $m$ minutes of simulated and experimental data with SRTM and Logan plot, where $m = [40, 45, 50, 55, 60, 65, 70, 75, 80, 85, 90, 95, 100, 105, 115, 120]$. For EQ, $BP_{ND}^{\text{Activation}}$ was taken as the average $BP_{ND}$ from 40 - $n$ minutes, where $n$ was [45, 50, 55, 60, 65, 70, 75, 80, 85, 90, 95, 100, 105, 110, 115, 120].

**Simplified reference tissue model (SRTM):** SRTM [1] assumes that tracer activity in both the ROI and reference region can be described by a 1-tissue compartment model. The three model parameters, $R_1$, $k_{12}$, and $k_{21}'$, were estimated by fitting SRTM to TACs using the following operational equation:

$$C_t(t) = R_1 C_{inv}(t) + R_1 (k_0 - k_{21}) C_{inv}(t) \otimes e^{-k_{21}' t}$$

(5)

where $C_t(t)$ is the activity of the tracer in the tissue, $C_{inv}(t)$ is the activity of the tracer in the reference region, $R_1 = k_{12}/k_{12}'$, the ratio of $k_1$ in the target tissue to $k_1'$ in the reference tissue, $k_{12}'$ is the efflux rate from the reference region, and $k_{21}'$ is the apparent efflux rate of tracer from the target region. The cerebellum was used as the reference region and cerebellar TACs were smoothed by fitting to a sum of exponentials. Parameters were estimated using unweighted least squares. Nonlinear parameter estimation was performed using a Marquardt-Levenberg algorithm [36] and custom software for IDL (8.0 ITT Visual Information Solutions, Boulder, CO). $BP_{ND}$ was calculated as:

$$BP_{ND} = R_1 (k_{12}')^{-1}$$

(6)

**Logan graphical method using a reference tissue as input:** The Logan graphical method using a reference tissue as input relies on the integration of tracer kinetic equations to yield a simple, linear equation whose slope equals the distribution volume ratio (DVR) of the tracer [2]:

$$\frac{\int C_t(u) du}{C_t(t)} = DVR \left( \frac{\int C_{ref}(u) du}{C_t(t)} \right) + \text{int}'$$

(7)

The Logan graphical technique proposes that after some time $t^*$, $\text{int}'$ is constant and the plot of $\frac{\int C_t(u) du}{C_t(t)}$ vs. $\frac{\int C_{ref}(u) du}{C_t(t)}$ becomes linear with slope DVR. In all Logan graphical method analyses, $t^*$ was fixed at 15 min and the cerebellum was used as the reference region. $BP_{ND}^{\text{EQ}}$ was then calculated as $DVR - 1$.

**Equilibrium analysis (EQ):** When tracer concentration is at equilibrium, $BP_{ND}$ is the ratio of the difference between target tissue concentration $C_{t, EQ}$ to reference tissue concentration $C_{ref, EQ}$ relative to the reference tissue concentration (Equation 8). $BP_{ND}$ is calculated as an average over multiple time frames:

$$BP_{ND} = \frac{C_{t, EQ} - C_{ref, EQ}}{C_{ref, EQ}}$$

(8)

where the superscript bar indicates average over some range of time frames. Equilibrium was achieved by 30 min in both simulated and experimental data.

**Results**

**Simulations**

**Effect of dopamine signal on simulated time activity curves:** Inclusion of a time-varying DA signal (Figure 2A) produced decreases in the simulated TACs dependent on the magnitude of the DA signal input (Figure 2B). Degree of deflection in TAC values following onset of DA release was significantly correlated ($R^2=0.96$, p < 0.005) with total DA concentration. Consistent with the transient nature of the DA inputs, activity level in many of the noiseless simulated TACs returned to baseline during the course of the scan (Figure 2B). In noisy simulations (Figure 2C), this effect was not easily visible with the naked eye.

**Effect of transient dopamine signal on SRTM fitting:** Fits of SRTM to simulated TACs contain-

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Fits to simulated TACs containing DA release had greater residual sum of squares than fits to rest TACs for activation scan window durations of 40 min or greater (the time of DA release start in simulations). After initiation of DA release, the shape of each fit became dependent on the duration of data used (Figure 3A). As the data window was widened to include more of the transient dip in the TAC, the fitted curves were pulled down relative to the simulated data. Fitted curves went back up again as the data window was widened even more to include later data points unaffected by the transient DA release. The same effect was seen when fitting SRTM to experimental data, though a deflection in experimental TACs at time of smoking was not easily visible (Figure 3B).

Effect of transient dopamine signal on Logan plot: Fits to the linear portion of the Logan plots were also dependent on the duration of data used. The fits varied with data window width and the size of the DA-induced changes in TACs. Subtle deviations from linearity, beginning at the onset of DA increase, were observed in the transformed data in Logan plot space for noiseless data (Figure 3C). However, the slopes of the transformed data before and after the time of DA release were the same. Deviations from linearity were not observed in the Logan plots of experimental data (Figure 3D).

Effect of scan duration on ΔBP

ΔBP varied based on the duration of data used to determine BP in the activation state for both noiseless (Figure 4) and noisy (Figure 5) simulations. This effect was seen for all DA signals tested. In simulations, a distinct pattern was seen: ΔBP increased as activation window duration increased, until reaching a maximum at some window width (generally 20-40 min after onset of DA release), after which ΔBP decreased with even wider activation window. The activation window width for which ΔBP was maximal varied by DA signal shape. Maximum ΔBP was found at longer activation windows for greater total DA release (Table 1). Note: greater total DA release was associated with later DA peak in our simulations. For a given DA signal, ΔBP values were highest when estimated by SRTM and lowest when estimated by the Logan graphical approach. For smaller DA responses (e.g., β=1), this resulted in negative ΔBP values for some activation window widths, an effect which was exacerbated by noise (Figure 5).

Rest data: All analyses were also performed on noiseless and noisy simulated Rest data (containing no DA release). ΔBP from Rest scan data (ΔBP Rest) did not vary noticeably by activation scan duration in any method tested (SRTM, Logan, or EQ) (Figure 6). In noisy simulations, there were biases in ΔBP Rest; for an activation window of 120 min ΔBP Rest was 0.8% ±

Figure 2. Dopamine (DA) signals with increasing area-under-the-curve (AUC) but same peak height (A) were used to simulate noiseless (B) and noisy (C) striatal time activity curves (TACs). All DA signals produced visually noticeable decreases in the noiseless simulated TACs. These decreases were less visually apparent for transient DA signals in the noisy simulated data (for ease of visualization, only one noisy curve is shown here). In (A), DA signals increase in width as β decreases.
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6.1% for SRTM, -4.1% ± 7.0% for the Logan graphical approach, and -1.0% ± 3.7% for EQ.

Human experiments

Δ\(BP_{ND}\) was observed to vary with activation scan duration in experimental data from 4 human subjects (Figure 7). A similar pattern to that seen from analysis of simulated data can be seen clearly in the left ventral striatum in the 2 male subjects and left dorsal caudate in the 2 female subjects (Figure 8). Dependence of Δ\(BP_{ND}\) on activation scan duration was not seen in other regions examined in these 4 subjects.

Discussion

This study examined the effect of the duration of the activation scan on Δ\(BP_{ND}\) estimated with SRTM, Logan plot, and EQ applied to PET data containing effects of transient DA release. Through analysis of simulations, we showed that the magnitude of Δ\(BP_{ND}\) is dependent on the amount of activation scan data used in estimation of \(BP_{ND, Activation}\). We found a similar effect in human data from a study of smoking-induced DA release. We believe that together, these findings illustrate the limitation of applying standard analysis methods to data acquired while the endogenous DA signal was varying rapidly during the scan.

The methods used in this study assume that \(BP_{ND}\) is a static parameter. For a method to be fruitful and reliable, varying the length of the scan data fitted should not change the \(BP_{ND}\) value estimated. As we show in our analysis of simulated Rest data (Figure 6), Δ\(BP_{ND, Rest}\) didn’t vary with activation scan duration. A slightly
biased $\Delta BP_{ND}^{\text{Rest}}$ was found by all methods and was greatest with the Logan plot. This suggests that more than 35 min of data may be necessary for these methods to produce unbiased $BP_{ND}$ estimates. In the case of the Logan plot, a longer $t^*$ for shorter scan durations (e.g. 35 min) may reduce this bias. Noise has also been shown to introduce a negative bias in $BP_{ND}$ estimated by the Logan plot [37].

Unlike simulated Rest data, $\Delta BP_{ND}$ varied with the activation scan duration for simulated data containing transient DA release. $\Delta BP_{ND}$ tended to decrease as more data after onset of DA release was included in the analysis (i.e. the data window was widened). Because the DA release in our simulations was transient and the effect of DA release on the tracer signal was brief, inclusion of data beyond the period

Figure 4. Effect of scan duration on $\Delta BP_{ND}$ estimated by SRTM (left column), Logan (middle column), and EQ (right column) for all noiseless simulated TACs. Each row represents the results from a noiseless simulated TAC containing the DA signal indicated by the $\beta$ value. For all graphs, the dependent axis is $\Delta BP_{ND}$ and the independent axis is activation window width in min.
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Table 1. Maximum $\Delta BP_{ND}$ values from noiseless simulation data

<table>
<thead>
<tr>
<th>DA signal</th>
<th>SRTM</th>
<th>Logan</th>
<th>EQ</th>
<th>SRTM</th>
<th>Logan</th>
<th>EQ</th>
<th>DA peak time</th>
</tr>
</thead>
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<tr>
<td>$\beta=2$</td>
<td>3%</td>
<td>1%</td>
<td>2%</td>
<td>60</td>
<td>55</td>
<td>40-50</td>
<td>41</td>
</tr>
<tr>
<td>$\beta=1$</td>
<td>5%</td>
<td>3%</td>
<td>4%</td>
<td>60</td>
<td>60</td>
<td>40-55</td>
<td>42</td>
</tr>
<tr>
<td>$\beta=0.5$</td>
<td>8%</td>
<td>6%</td>
<td>7%</td>
<td>65</td>
<td>65</td>
<td>40-60</td>
<td>44</td>
</tr>
<tr>
<td>$\beta=0.25$</td>
<td>12%</td>
<td>10%</td>
<td>11%</td>
<td>75</td>
<td>75</td>
<td>40-75</td>
<td>48</td>
</tr>
<tr>
<td>$\beta=0.15$</td>
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<td>14%</td>
<td>14%</td>
<td>90</td>
<td>90</td>
<td>40-85</td>
<td>53</td>
</tr>
</tbody>
</table>

Figure 5. Effect of scan duration on $\Delta BP_{ND}$ estimated by SRTM (left column), Logan (middle column), and EQ (right column) for all noisy simulated TACs. Points are mean values and error bars are SD of 100 noisy curves. Each row represents the results from noisy simulated TACs containing the DA signal indicated by the $\beta$ value. For all graphs, the dependent axis is $\Delta BP_{ND}$ and the independent axis is activation window width in min.

of DA release appears to have washed out the effect of the DA release on $\Delta BP_{ND}$. This suggests that at typical scan durations (90 min-120 min) used in $[^{11}C]$raclopride studies, $\Delta BP_{ND}$
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measured by these common methods may have low sensitivity to transient DA release.

Endres and Carson [22] showed that when neurotransmitter release is time-varying, distribution volume (and hence binding potential) $\Delta BP_{ND}$ is a time-weighted average of $\Delta BP_{ND}(t)$ weighted by the free tracer concentration [22]. Our simulations are a demonstration of this principle. We also show that if DA release is transient and total DA release is modest, it may not be detectable via $\Delta BP_{ND}$ depending on the duration of activation scan data analyzed. This is due not only to the kinetics of endogenous DA but also those of the tracer. Yoder et al. [38] showed that $\Delta BP_{ND}$ is dependent on the relative timing and kinetics of both the tracer and the endogenous ligand. In our simulation studies, onset of DA release was 40 min after time of injection. Had DA release occurred earlier during the activation scan, $\Delta BP_{ND}$ would be even smaller for longer (90, 120 min) scan durations because there would be more PET data which had returned to baseline levels after DA release included in the estimation of $BP_{ND}$. This would further decrease the likelihood of detection of transient DA release by $\Delta BP_{ND}$ for typical $[^{11}C]raclopride$ scan durations of 90 or 120 min.

Our simulations suggest that the discrepancy in $\Delta BP_{ND}$ between findings in two separate smoking studies by Brody et al. [11, 13] may be due solely to differences in the amount of data used to estimate $BP_{ND}^{\text{Activation}}$. In the studies by Brody et al., subjects were scanned for 50 min,
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taken off the camera for 10 min with the tracer infusion still running and allowed to smoke 1 cigarette outside, then returned to the camera and scanned for an additional 30 min. EQ was used to estimate \( \text{BP}_{\text{ND}} \). In an early study [13], \( \text{BP}_{\text{ND, Activation}} \) was estimated from the average of 10 min of data collected after each subject smoked a cigarette (60-70 min post-injection), and a 30% decrease in \( \text{BP}_{\text{ND}} \) due to smoking was reported in the left ventral caudate. In later studies [9-12], 30 min of data (60-90 min post-injection) were used to estimate \( \text{BP}_{\text{ND, Activation}} \), and only an 8% decrease in \( \text{BP}_{\text{ND}} \) was found in the left ventral caudate. Brody et al. did not offer a persuasive argument to explain this large discrepancy between studies. However, the drop in effect size with increased activation scan window is consistent with our simulation findings which show that for \( \beta = 1 \), an activation scan window of 40-60 min gives \( \Delta \text{BP}_{\text{ND}} = 4\% \) but an activation scan window of 40-80 min gives \( \Delta \text{BP}_{\text{ND}} = 2\% \) (Figure 4). The actual \( \Delta \text{BP}_{\text{ND}} \) from the Brody et al. studies depends on how big the DA effect was and its temporal kinetics, as well as other factors such as how the regions were drawn, but our simulations predict that there would be a discrepancy in \( \Delta \text{BP}_{\text{ND}} \) reported using 20 min vs. 40 min of activation scan data.

Figure 8. Effect of scan duration on \( \Delta \text{BP}_{\text{ND}} \) estimated by SRTM (left column), Logan (middle column), and EQ (right column) for experimental data. Arrows indicate times at which cigarettes were smoked. \( \Delta \text{BP}_{\text{ND}} \) values estimated from human smoking studies were dependent on the activation scan window duration. The pattern seen in simulations was identified in the left ventral striatum (top two rows) of two male subjects (M1 and M2) and in the left dorsal caudate (bottom two rows) of two female subjects (F1 and F2). Note: Unlike subjects M1, M2, and F1, who smoked 2 cigarettes at 45 min after time of injection (TOI), subject F2 smoked two cigarettes ~30 min apart, at 53 min and 87 min after TOI.
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precisely in the direction of the observed discrepancy in the Brody et al. studies.

The dependence of $\Delta BP_{ND}$ on activation scan duration demonstrated by simulations was also observed in analysis of our human data. This pattern seen in simulations was seen only in the left ventral striatum of the two male subjects studies and only in the left dorsal caudate of the two female subjects. Based on our simulations, this suggests the presence of transient DA release in these regions. The regional difference in DA release in males and females seen here may suggest a gender-effect in DA response to smoking, one which we are currently pursuing.

The interpretation of these results is limited to the framework in which these studies were completed. We tested SRTM, the Logan graphical method, and EQ on B/I data from a single-scan activation study design. While application of EQ to this study design is common, SRTM and the Logan graphical method are typically applied to data from dual-scan (aka “paired-bolus”) activation studies. Nonetheless, we believe that application of the methods as done here demonstrates a general danger of applying conventional analysis methods to $[11C]$raclopride data which contain transient DA release. When the data violate the assumption that DA concentration is at steady-state, $BP_{ND}$ is no longer a static parameter and the measured $\Delta BP_{ND}$ becomes a time-weighted average of the instantaneous $\Delta BP_{ND}(t)$. Practically speaking, this may increase inter-subject variability in $\Delta BP_{ND}$ since the kinetics of DA release may vary across subjects. This variability will, in turn, reduce effect size.

This would seem to suggest that shorter scan durations may improve the detection of transient DA release via $\Delta BP_{ND}$. However, because $\Delta BP_{ND}$ depends on the timing and kinetics of DA relative to that of the tracer, interpretation of $\Delta BP_{ND}$ as an index of the amount of DA released is not straightforward. Instead of optimizing conventional methods to detect a significant $\Delta BP_{ND}$, it may be better to estimate DA release using advanced models, which take into account the time-variation in DA signal. The linear extension of the simplified reference region model (LSRRM) [39] and the neurotransmitter PET (ntPET) suite of methods [40-43] both allow endogenous neurotransmitter to vary during the course of the scan. These methods may be better suited for estimating transient DA release, such as smoking-induced DA release, from $[11C]$raclopride PET data.

Conclusions

The dependence of $\Delta BP_{ND}$, as estimated by SRTM, Logan plot, and EQ method, on the duration of activation data analyzed illustrates the inability of conventional methods based on time-invariant parameters to reliably quantify short-lived increases in endogenous DA. Ultimately, more complicated models which take into account time-variation in DA are needed to properly measure the effect of stimulus-induced DA release when the DA response is brief.

Acknowledgments

We'd like to thank Drs. Karmen Yoder and Dianne Lee for their assistance with defining our striatal template. We thank Erin McGovern and Sabrina Helmbrecht for subject recruitment and screening. We also acknowledge the chemistry staff, nuclear medicine technologists, nurses, and research staff of the Yale PET Center without whom this work would not be possible.

Supported by ORWH and NIDA (P50DA033945, K02 DA031750, R21DA032791).

Conflict of interest statement

The authors have no conflicts of interest to declare.

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